

Variational formulations of Bayesian Inference

Overview of materials

Lec 1 1) Review of Bayesian (finite dimensional)
parametric inference

2) Infinite dimensional Bayesian
inference - Gaussian process
model

Lec 2 3) Variational formulations
of gradient updating
and gradient flows

Lec 3 4) Posterior consistency
a) Classic test fn/
Sieve approach
b) Variational approach

based on large
deviations

Lec 4 5) Bayesian inverse
problems

Bayes : Two manuscripts

1) Bayes' rule

2) A response to :

The analyst: or, a discourse
addressed to an infidel
mathematician

Given sets A & B

$$P(B|A) = \frac{P(A|B) P(B)}{P(A)}$$

there is a symmetry

Bayesian inference:

Likelihood or data generating process $f(x_1, \dots, x_n | \theta)$, $\theta \in \Theta$

$$f(x_1, \dots, x_n) = \prod_{i=1}^n f(x_i | \theta) \quad \text{iid}$$

Prior: $\pi(\theta)$ - belief you see data

Posterior:

$$\pi(\theta | x_1, \dots, x_n) = \frac{f(x_1, \dots, x_n | \theta) \pi(\theta)}{\int_{\Theta} f(x_1, \dots, x_n | \theta) \pi(\theta) d\theta}$$

||
marginal likelihood
or $P(D)$

There is an asymmetry.

- a) Quantifies uncertainty
- b) Need to specify a likelihood
- c) Need a prior
- d) The normalization constant

Contrast to frequentist/proceduralist approach.

Ex. 1. Binomial

$$f(x | n, p) = \binom{n}{x} p^x (1-p)^{n-x}$$

$$\pi(p) \sim \text{Beta}(\alpha, \beta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha) \Gamma(\beta)} p^{\alpha-1} (1-p)^{\beta-1}$$

$\alpha, \beta > 0$

$$\pi(p | x, \alpha, \beta) = \frac{p^{x+\alpha-1} (1-p)^{n-x+\beta+1}}{\int_0^1 p^{x+\alpha-1} (1-p)^{n-x+\beta+1} dp}$$

$$= \text{Beta}(x+1, n-x+1, \beta)$$

The Beta is conjugate prior for the binomial likelihood.

Posterior consistency: The posterior $\pi_n(\theta | x_1, \dots, x_n)$ is consistent at θ_0 if for every neighborhood U of θ_0

$$\pi_n(U) \rightarrow 1 \text{ as } n \rightarrow \infty \text{ under } \theta_0.$$

For observations with finite # of values at any point θ_0 which belongs in the support of π .

Counterexample:

Infinite multinomial - Infer a pmf on the set of integers.

Let $\theta_0: \Pr(X=k) = (1-p)^{k-1} p$
be geometric.

One can construct a prior π that gives positive mass to every neighborhood of θ_0 but the posterior concentrates around

$$\theta_n: \Pr(X=k) = (1-p)^{k-1} p$$

Bayesian hierarchical model:

Often count data is modeled as λ Poisson

$$f(m|\lambda) = \frac{\lambda^m e^{-\lambda}}{m!}$$

$$E(m) = \text{Var}(m) = \lambda$$

Often $\text{Var}(m) > E(m)$

$$\lambda \sim \text{Gamma}(r, \frac{p}{1-p})$$

$$f(\lambda | r, p) = \frac{\left(\frac{p}{1-p}\right)^r}{\Gamma(r)} \lambda^{r-1} e^{-\lambda \frac{p}{1-p}}$$

$$m | \lambda \sim \text{Pois}(\lambda)$$

$$f(m | \lambda) = \frac{e^{-\lambda} \lambda^m}{m!}$$

$$f(m, \lambda | r, p) = \frac{\left(\frac{p}{1-p}\right)^r}{\Gamma(r)} \lambda^{r-1} e^{-\lambda \frac{p}{1-p}} \frac{e^{-\lambda} \lambda^m}{m!}$$

Integrate out λ :

$$\int_0^{\infty} f(m, \lambda | r, p) d\lambda =$$

$$f(m | r, p) = \frac{\Gamma(m+r)}{\Gamma(r) m!} p^r (1-p)^m$$

of r successes before m failures

Exponential family:

Given a family of distributions with real parameters, $\theta = [\theta_1, \dots, \theta_k]^T$

The family can be written

$$f(x|\theta) = h(x) g(\theta) \exp(\eta(\theta) \cdot T(x))$$

T = sufficient statistic

η = natural parameterization.

$$f(y; \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(y-\mu)^2}{2\sigma^2}}$$

$$\eta = \left[\frac{\mu}{\sigma^2} \quad -\frac{1}{2\sigma^2} \right]$$

$$h(y) = \frac{1}{\sqrt{2\pi}}$$

$$T(y) = (y, y^2)^T$$

$$g(\theta) = \frac{\mu^2}{2\sigma^2} + \log|\sigma|$$

Diaconis & Ylvisaker:

Given an exponential family likelihood conjugate priors satisfy

$$\mathbb{E}(\mathbb{E}(x|a) | X=x) = ax + b$$

Gaussian process - Infinite dimensional model.

Defn. A Gaussian process $\{X_t\}_{t \in T}$ indexed by a set T is a family of r.v.'s that for any finite subset $F \subset T$ $X_F := \{X_t\}_{t \in F}$ is MVN. If X_F is non-degenerate MVT for all F then $\{X_t\}$ is a non-degenerate GP.

For GP $\{X_t\}_{t \in T}$ with mean fun.

$$\mu_t = \mathbb{E}(X_t)$$

and covariance kernel ^{is positive}
 $R(X_s, X_t) = \text{Cov}(X_s, X_t)$ ^{(semi) definite}

and $\{X_t\}$ is non-degenerate
then for any finite FCT

$$X_F \sim \text{MVN}(\mu_F, \Sigma_F)$$

$$\mu_F = \begin{pmatrix} \mu(X_1) \\ \vdots \\ \mu(X_{F_n}) \end{pmatrix}$$

$$\Sigma_{F,i,j} = \text{Cov}(X_i, X_j)$$

GP examples:

Brownian motion - W_t

$$\mathbb{E}(W_s W_t) = \min(s, t)$$

Ornstein-Uhlenbeck - Y_t

$$\mathbb{E}(Y_s Y_t) = \exp(-|t-s|)$$

Brownian bridge - W_t°

$$\mathbb{E}(W_s^\circ W_t^\circ) = \min(s, t) - st$$

Some theory and practice of GP

GP prior regression

Prior - $f \sim \text{GP}(\mu(\cdot), K(\cdot, \cdot))$

$$\mu(x) = \mathbb{E}(f(x))$$

$$K(x_i, x_j) = \text{Cov}(f(x_i), f(x_j))$$

Likelihood model:

$$Y_i = f(x_i) + \varepsilon_i, \\ \varepsilon_i \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$$

$$X = \begin{bmatrix} -x_1- \\ -x_2- \\ \vdots \\ -x_n- \end{bmatrix}, \quad X^* = \begin{bmatrix} -x_1^* - \\ \vdots \\ -x_n^* - \end{bmatrix},$$

$$Y = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}, \quad Y^* = \begin{bmatrix} y_1^* \\ \vdots \\ y_n^* \end{bmatrix}$$

$$\varepsilon = \begin{bmatrix} \varepsilon_1 \\ \vdots \\ \varepsilon_n \end{bmatrix}, \quad \varepsilon^* = \begin{bmatrix} \varepsilon_1^* \\ \vdots \\ \varepsilon_n^* \end{bmatrix}$$

$$f = \begin{bmatrix} f(x_1) \\ \vdots \\ f(x_n) \end{bmatrix}, f^* = \begin{bmatrix} f^*(x_1) \\ \vdots \\ f^*(x_{n^*}) \end{bmatrix}$$

want $Y^* | X, X^* \sim N(\mu^*, \Sigma^*)$

$$\begin{bmatrix} Y \\ Y^* \end{bmatrix} | X^*, X = \begin{bmatrix} f \\ f^* \end{bmatrix} + \begin{bmatrix} \varepsilon \\ \varepsilon^* \end{bmatrix}$$

$$\sim N\left(0, \begin{bmatrix} K(x, x) + \sigma^2 I & K(x^*, x) \\ K(x^*, x) & K(x^*, x^*) + \sigma^2 I \end{bmatrix}\right)$$

$$\mu^* = K(x^*, x) [K(x, x) + \sigma^2 I]^{-1} Y$$

$$\Sigma^* = K(x^*, x^*) + \sigma^2 I - \\ K(x^*, x) [K(x, x) + \sigma^2 I]^{-1} K(x, x^*)$$

Some theory - Empirical

Process theory

1) GPs on manifolds

Given compact Riemannian manifold (M, g) can we define a GP supported on M ?

Näive idea.

$$K(x, x') = \sigma^2 \exp\left(-\frac{d_g(x, x')^2}{2\kappa^2}\right) \quad (5)$$

Theorem: If (5) is PSD for all $\kappa^2 > 0$ then \mathcal{R} is isometric to a Euclidean space.

Solution:

$$\left(\frac{2\nu}{\kappa^2} - \Delta_M\right)^{\frac{\nu}{2} + \frac{d}{4}} f = \mathcal{W}_\nu$$

$$e^{-\frac{\nu}{2} \Delta_M} f = \mathcal{W}$$

to \Uparrow solve SPDEs
get $K(x, x')$

2) Extrema of Gaussian processes

Def. 2.1: Let (T, d) be a compact metric space. For each $\varepsilon > 0$ the (Lebesgue) covering # $N(\varepsilon)$ is the minimum number of ε -balls to cover T .

Thm: Let d be the canonical metric

$$d(s, t) = \sqrt{|\Sigma| |X_s - X_t|^2}$$

of a non-degenerate, centered, GP an $N(\varepsilon)$ is the covering #.

If for some $\rho > 0$

$$\int_0^\rho \sqrt{\log N(\epsilon)} d\epsilon < \infty$$

the GP has uniformly continuous sample paths
and

$$\sup_{t \in T} X_t = \max_{t \in T} X_t < \infty.$$

Concentration of sup:

Thm: Let $\{X_t\}_{t \in T}$ be a centered GP on a countable T that is a.s. bounded. $X^* = \sup_t X_t$.

If $\sigma_T^2 := \sup_{t \in T} \mathbb{E} X_t^2 < \infty$,
then $\mathbb{E} X^* < \infty$ and for any $u > 0$

$$P(X^* \geq \mathbb{E} X^* + u) \leq e^{-u^2 / 2\sigma_T^2}$$

3) Reproducing Kernel Hilbert space

Given a positive definite kernel $k(\cdot, \cdot)$ that is

$\forall t_1, \dots, t_n \in X$ + $\forall a_1, \dots, a_n \in \mathbb{R}$
and all $n \in \mathbb{N}$

$$\sum_{i,j}^n a_i a_j k(x_i, x_j) > 0.$$

Define $L_k : L_2(X) \rightarrow C(X)$

$$L_k f := \int_X k(s, t) f(t) dt = g(t)$$

Eigenvalues + eigenvectors

$$\int_X k(s, t) \varphi_k(t) dt = \lambda_k \varphi_k(t)$$

$\forall k$

$$k(s, t) = \sum_j \lambda_j \varphi_j(s) \varphi_j(t)$$

$$\|f(s)\|_{H_k}^2 = \sum_j \langle c_j \varphi_j(s), c_j \varphi_j(s) \rangle$$

~ . . .

$$:= \sum_j \frac{c_j}{\lambda_j}$$

$$\begin{aligned} \langle f, g \rangle &= \left\langle \sum_j c_j \vartheta_j(s), \sum_j d_j \vartheta_j(s) \right\rangle \\ &= \sum_j \frac{c_j d_j}{\lambda_j} \end{aligned}$$

$$H_k = \left\{ \underbrace{f \mid f(s) = \sum_k c_k \vartheta_k(s)}_{\text{and } \|f\|_{H_k} < \infty} \right\}$$

$$\langle f(\cdot), k(\cdot, x) \rangle_{H_k} = f(x)$$